

# Tolerance Limits on Order Statistics in Future Samples Coming from the Two-Parameter Exponential Distribution

Nicholas A. Nechval<sup>1, \*</sup>, Konstantin N. Nechval<sup>2</sup>

<sup>1</sup>Department of Mathematics, Baltic International Academy, Riga, Latvia

<sup>2</sup>Department of Applied Mathematics, Transport and Telecommunication Institute, Riga, Latvia

## Email address:

nechval@junik.lv (N. A. Nechval), konstan@tsi.lv (K. N. Nechval)

## To cite this article:

Nicholas A. Nechval, Konstantin N. Nechval. Tolerance Limits on Order Statistics in Future Samples Coming from the Two-Parameter Exponential Distribution. *American Journal of Theoretical and Applied Statistics*. Special Issue: Novel Ideas for Efficient Optimization of Statistical Decisions and Predictive Inferences under Parametric Uncertainty of Underlying Models with Applications. Vol. 5, No. 2-1, 2016, pp. 1-6. doi: 10.11648/j.ajtas.s.2016050201.11

**Abstract:** This paper presents an innovative approach to constructing lower and upper tolerance limits on order statistics in future samples. Attention is restricted to invariant families of distributions under parametric uncertainty. The approach used here emphasizes pivotal quantities relevant for obtaining tolerance factors and is applicable whenever the statistical problem is invariant under a group of transformations that acts transitively on the parameter space. It does not require the construction of any tables and is applicable whether the past data are complete or Type II censored. The proposed approach requires a quantile of the  $F$  distribution and is conceptually simple and easy to use. For illustration, the two-parameter exponential distribution is considered. A practical example is given.

**Keywords:** Order Statistics,  $F$  Distribution, Lower Tolerance Limit, Upper Tolerance Limit

## 1. Introduction

Statistical tolerance limits are another tool for making statistical inference on an unknown population. As opposed to a confidence limit that provides information concerning an unknown population parameter, a tolerance limit provides information on the entire population; to be specific, one-sided tolerance limit is expected to capture a certain proportion or more of the population, with a given confidence level. For example, an upper tolerance limit for a univariate population is such that with a given confidence level, a specified proportion or more of the population will fall below the limit. A lower tolerance limit satisfies similar conditions.

It is often desirable to have statistical tolerance limits available for the distributions used to describe time-to-failure data in reliability problems. For example, one might wish to know if at least a certain proportion, say  $\beta$ , of a manufactured product will operate at least  $T$  hours. This question can not usually be answered exactly, but it may be possible to determine a lower tolerance limit  $L(X_1, \dots, X_n)$ , based on a preliminary random sample  $(X_1, \dots, X_n)$ , such that one can say with a certain confidence  $\gamma$  that at least  $100\beta\%$  of the product will operate longer than  $L(X_1, \dots, X_n)$ . Then reliability

statements can be made based on  $L(X_1, \dots, X_n)$ , or, decisions can be reached by comparing  $L(X_1, \dots, X_n)$  to  $T$ . Tolerance limits of the type mentioned above are considered in this paper. That is, if  $f_\theta(x)$  denotes the density function of the parent population under consideration and if  $S$  is any statistic obtained from the preliminary random sample  $(X_1, \dots, X_n)$  of that population, then  $L(S)$  is a lower  $\gamma$  probability tolerance limit for proportion  $\beta$  if

$$\Pr \left( \int_{L(S)}^{\infty} f_\theta(x) dx \geq \beta \right) = \gamma, \quad (1)$$

and  $U(S)$  is an upper  $\gamma$  probability tolerance limit for proportion  $\beta$  if

$$\Pr \left( \int_{-\infty}^{U(S)} f_\theta(x) dx \geq \beta \right) = \gamma, \quad (2)$$

where  $\theta$  is the parameter (in general, vector).

The common distributions used in life testing problems are the normal, exponential, Weibull, and gamma distributions [1]. Tolerance limits for the normal distribution have been

considered in [2], [3], [4], and others.

Tolerance limits enjoy a fairly rich history in the literature and have a very important role in engineering and manufacturing applications. Patel [5] provides a review (which was fairly comprehensive at the time of publication) of tolerance intervals for many distributions as well as a discussion of their relation with confidence intervals for percentiles and prediction intervals. Dunsmore [6] and Guenther, Patil, and Uppuluri [7] both discuss 2-parameter exponential tolerance intervals and the estimation procedure in greater detail. Engelhardt and Bain [8] discuss how to modify the formulas when dealing with type II censored data. Guenther [9] and Hahn and Meeker [10] discuss how one-sided tolerance limits can be used to obtain approximate two-sided tolerance intervals by applying Bonferroni's inequality.

In contrast to other statistical limits commonly used for statistical inference, the tolerance limits (especially for the order statistics) are used relatively rarely. One reason is that the theoretical concept and computational complexity of the tolerance limits is significantly more difficult than that of the standard confidence and prediction limits. Thus it becomes necessary to use the innovative approaches which will allow one to construct tolerance limits on future order statistics for many populations.

In this paper, the innovative approach to constructing lower and upper tolerance limits on order statistics in future samples is proposed. For illustration, the two-parameter exponential distribution is considered.

## 2. Mathematical Preliminaries

### 2.1. Probability Distribution Function of Order Statistic

*Theorem 1.* If there is a random sample of  $m$  ordered observations  $Y_1 \leq \dots \leq Y_m$  from a known distribution (continuous or discrete) with density function  $f_\theta(y)$ , distribution function  $F_\theta(y)$ , then the probability distribution function of the  $k$ th order statistic  $Y_k$ ,  $k \in \{1, 2, \dots, m\}$ , is given by

$$\begin{aligned} P_\theta(Y_k \leq y_k | m) &= \sum_{j=k}^m \binom{m}{j} [F_\theta(y_k)]^j [1 - F_\theta(y_k)]^{m-j} \\ &= \int_{\frac{1-F_\theta(y_k)}{F_\theta(y_k)} \frac{2k}{2(m-k+1)}}^{\infty} f_{2(m-k+1), 2k}(x) dx, \end{aligned} \quad (3)$$

where

$$\begin{aligned} f_{2(m-k+1), 2k}(x) &= \frac{1}{\text{B}\left(\frac{2(m-k+1)}{2}, \frac{2k}{2}\right)} \\ &\times \left(\frac{2(m-k+1)}{2k}\right) \left(\frac{2(m-k+1)}{2k}x\right)^{2(m-k+1)/2-1} \end{aligned}$$

$$\times \left(1 + \frac{2(m-k+1)}{2k}x\right)^{-[2(m-k+1)+2k]/2}, \quad x > 0, \quad (4)$$

is the probability density function of an  $F$  distribution with  $2(m-k+1)$  and  $2k$  degrees of freedom.

*Proof.* Suppose an event occurs with probability  $p$  per trial. It is well-known that the probability  $P$  of its occurring  $k$  or more times in  $m$  trials is termed a cumulative binomial probability, and is related to the incomplete beta function  $I_x(a, b)$  as follows:

$$P \equiv \sum_{j=k}^m \binom{m}{j} p^j (1-p)^{m-j} = I_p(k, m-k+1). \quad (5)$$

It follows from (5) that

$$\begin{aligned} P_\theta\{Y_k \leq y_k | m\} &= \sum_{j=k}^m \binom{m}{j} [F_\theta(y_k)]^j [1 - F_\theta(y_k)]^{m-j} \\ &= I_{F_\theta(y_k)}(k, m-k+1) \\ &= \frac{1}{\text{B}(k, m-k+1)} \int_0^{F_\theta(y_k)} u^{k-1} (1-u)^{(m-k+1)-1} du \\ &= \frac{\left(\frac{2(m-k+1)}{2k}\right)^{2(m-k+1)/2}}{\text{B}\left(\frac{2k}{2}, \frac{2(m-k+1)}{2}\right)} \int_0^{F_\theta(y_k)} u^{\frac{2(m-k+1)+2k}{2}} \\ &\times \left(\frac{1-u}{u} \frac{2k}{2(m-k+1)}\right)^{2(m-k+1)/2-1} \frac{-2k}{2(m-k+1)} \left(\frac{du}{u^2}\right) \\ &= \frac{\left(\frac{2(m-k+1)}{2k}\right)^{2(m-k+1)/2}}{\text{B}\left(\frac{2(m-k+1)}{2}, \frac{2k}{2}\right)} \int_{\frac{1-F_\theta(y_k)}{F_\theta(y_k)} \frac{2k}{2(m-k+1)}}^{\infty} x^{2(m-k+1)/2-1} \\ &\times \left(1 + \frac{2(m-k+1)}{2k}x\right)^{-[2(m-k+1)+2k]/2} dx, \end{aligned} \quad (6)$$

where

$$x = \frac{1-u}{u} \frac{2k}{2(m-k+1)} \quad (7)$$

This ends the proof.

*Corollary 1.1.*

$$P_\theta(Y_k > y_k | m) = \sum_{j=0}^{k-1} \binom{m}{j} [F_\theta(y_k)]^j [1 - F_\theta(y_k)]^{m-j}$$

$$= \int_0^{\frac{1-F_\theta(y_k)}{F_\theta(y_k)} \frac{2k}{2(m-k+1)}} f_{2(m-k+1),2k}(x) dx. \tag{8}$$

Corollary 1.2. If  $y_{k,m;\gamma}$  is the quantile of order  $\gamma$  for the distribution of  $Y_k$ , we have from (3) that  $y_{k,m;\gamma}$  is the solution of

$$F_\theta(y_{k,m;\gamma}) = \frac{k}{k + (m - k + 1)q_{2(m-k+1),2k;1-\gamma}}, \tag{9}$$

where  $q_{2(m-k+1),2k;1-\gamma}$  is the quantile of order  $1-\gamma$  for the  $F$  distribution with  $2(m-k+1)$  and  $2k$  degrees of freedom.

### 2.2. Two-Parameter Exponential Distribution

The two-parameter exponential distribution is a widely used and widely known distribution. It is characterized by the density function

$$f_\theta(y) = \frac{1}{\sigma} \exp\left(-\frac{y - \mu}{\sigma}\right), \quad y \geq \mu, \tag{10}$$

where  $\theta = (\mu, \sigma)$ ,  $\sigma$  (scale parameter) and  $\mu$  (shift parameter) are unknown. The distribution function of the two-parameter exponential distribution is

$$F_\theta(y) = 1 - \exp\left(-\frac{y - \mu}{\sigma}\right) \tag{11}$$

The sufficient statistic for the parameter  $\theta$ , based on the  $r$  ( $\leq n$ ) smallest observations ( $X_1 \leq \dots \leq X_r$ ) in a random sample of size  $n$  from the two-parameter exponential distribution (10) is

$$S = \left( X_1 = \min(X_1, \dots, X_r), S_1 = \sum_{i=1}^r X_i + (n - r)X_r \right), \tag{12}$$

$$X_1 \sim h_\theta(x_1) = \frac{n}{\sigma} \exp\left(-\frac{n(x_1 - \mu)}{\sigma}\right), \quad x_1 \geq \mu, \tag{13}$$

where

$$V = \frac{X_1 - \mu}{\sigma} \tag{14}$$

is the pivotal quantity with the density function

$$h(v) = n \exp(-nv), \quad v \geq 0, \tag{15}$$

$$S_1 \sim g_\sigma(s_1) = \frac{1}{\Gamma(r-1)\sigma^{r-1}} s_1^{r-2} \exp\left(-\frac{s_1}{\sigma}\right), \quad s_1 \geq 0, \tag{16}$$

where

$$W = \frac{S_1}{\sigma} \tag{17}$$

is the pivotal quantity with the density function

$$g(w) = \frac{1}{\Gamma(r-1)} w^{r-2} \exp(-w), \quad w \geq 0. \tag{18}$$

## 3. Tolerance Limits for Order Statistic

### 3.1. Lower Tolerance Limit

Theorem 2. Let  $X_1 \leq \dots \leq X_r$  be the first  $r$  ordered observations from the preliminary sample of size  $n$  from a two-parameter exponential distribution defined by the density function (10). Then a lower one-sided  $\beta$ -content tolerance limit at level  $\gamma$ ,  $L_k \equiv L_k(S)$  (on the  $k$ th order statistic  $Y_k$  from a set of  $m$  future ordered observations  $Y_1 \leq \dots \leq Y_m$  also from the distribution (10)), which satisfies

$$\Pr(P_\theta(Y_k > L_k | m) \geq \beta) = \gamma, \tag{19}$$

is given by

$$L_k = \begin{cases} X_1 + \frac{S_1}{n} \left[ 1 - \left( \frac{\delta_\beta^n}{1 - \gamma} \right)^{\frac{1}{r-1}} \right], & \text{if } n \geq \frac{\ln(1 - \gamma)}{\ln \delta_\beta}, \\ X_1 - \frac{S_1}{n} \left[ \left( \frac{\delta_\beta^n}{1 - \gamma} \right)^{\frac{1}{r-1}} - 1 \right], & \text{if } n < \frac{\ln(1 - \gamma)}{\ln \delta_\beta}, \end{cases} \tag{20}$$

where

$$\delta_\beta = \frac{(m - k + 1)q_{2(m-k+1),2k;\beta}}{(m - k + 1)q_{2(m-k+1),2k;\beta} + k}. \tag{21}$$

Proof. It follows from (8), (11) and (19) that

$$\begin{aligned} & \Pr(P_\theta(Y_k > L_k | m) \geq \beta) \\ &= \Pr\left( \int_0^{\frac{1-F_\theta(L_k)}{F_\theta(L_k)} \frac{2k}{2(m-k+1)}} f_{2(m-k+1),2k}(x) dx \geq \beta \right) \\ &= \Pr\left( \frac{1 - F_\theta(L_k)}{F_\theta(L_k)} \frac{2k}{2(m - k + 1)} \geq q_{2(m-k+1),2k;\beta} \right) \\ &= \Pr\left( F_\theta(L_k) \leq \frac{k}{k + (m - k + 1)q_{2(m-k+1),2k;\beta}} \right) \\ &= \Pr\left( \exp\left(-\frac{L_k - \mu}{\sigma}\right) \geq \frac{(m - k + 1)q_{2(m-k+1),2k;\beta}}{(m - k + 1)q_{2(m-k+1),2k;\beta} + k} \right) \end{aligned}$$

$$\begin{aligned}
&= \Pr\left(\frac{L_k - \mu}{\sigma} \leq -\ln\left(\frac{(m-k+1)q_{2(m-k+1),2k;\beta}}{(m-k+1)q_{2(m-k+1),2k;\beta} + k}\right)\right) \\
&= \Pr\left(V \leq -\eta_L W - \ln \delta_\beta\right) = \int_0^{-\eta_L w - \ln \delta_\beta} h(v)dv, \quad (22)
\end{aligned}$$

where

$$\eta_L = \frac{L_k - X_1}{S_1}, \quad (23)$$

is the lower tolerance factor,

$$\delta_\beta = \frac{(m-k+1)q_{2(m-k+1),2k;\beta}}{(m-k+1)q_{2(m-k+1),2k;\beta} + k}. \quad (24)$$

It follows from (18), (19) and (22) that

$$\begin{aligned}
\eta_L &= \arg\left(\int_0^\infty \int_0^{-\eta_L w - \ln \delta_\beta} h(v)g(w)dv dw = \gamma\right) \\
&= \frac{1}{n} \left[ 1 - \left( \frac{\delta_\beta^n}{1-\gamma} \right)^{1/(r-1)} \right]. \quad (25)
\end{aligned}$$

Taking into account (23) and (25), we have (20). This completes the proof.

*Corollary 2.1.* If  $k=m=1$ , then

$$L_k = \begin{cases} X_1 + \frac{S_1}{n} \left[ 1 - \left( \frac{\beta^n}{1-\gamma} \right)^{\frac{1}{r-1}} \right], & \text{if } n \geq \frac{\ln(1-\gamma)}{\ln \beta}, \\ X_1 - \frac{S_1}{n} \left[ \left( \frac{\beta^n}{1-\gamma} \right)^{\frac{1}{r-1}} - 1 \right], & \text{if } n < \frac{\ln(1-\gamma)}{\ln \beta}. \end{cases} \quad (26)$$

The result similar to that of (26) can be found in [7, 8].

### 3.2. Upper Tolerance Limit

*Theorem 3.* Let  $X_1 \leq \dots \leq X_r$  be the first  $r$  ordered observations from the preliminary sample of size  $n$  from a two-parameter exponential distribution defined by the density function (10). Then a lower one-sided  $\beta$ -content tolerance limit at level  $\gamma$ ,  $U_k \equiv U_k(S)$  (on the  $k$ th order statistic  $Y_k$  from a set of  $m$  future ordered observations  $Y_1 \leq \dots \leq Y_m$  also from the distribution (10)), which satisfies

$$\Pr(P_\theta(Y_k \leq U_k | m) \geq \beta) = \gamma, \quad (27)$$

is given by

$$U_k = \begin{cases} X_1 + \frac{S_1}{n} \left[ 1 - \left( \frac{\delta_{1-\beta}^n}{\gamma} \right)^{\frac{1}{r-1}} \right], & \text{if } n \geq \frac{\ln \gamma}{\ln \delta_{1-\beta}}, \\ X_1 - \frac{S_1}{n} \left[ \left( \frac{\delta_{1-\beta}^n}{\gamma} \right)^{\frac{1}{r-1}} - 1 \right], & \text{if } n < \frac{\ln \gamma}{\ln \delta_{1-\beta}}, \end{cases} \quad (28)$$

where

$$\delta_{1-\beta} = \frac{(m-k+1)q_{2(m-k+1),2k;1-\beta}}{(m-k+1)q_{2(m-k+1),2k;1-\beta} + k}. \quad (29)$$

*Proof.* It follows from (3), (11) and (27) that

$$\begin{aligned}
&\Pr(P_\theta(Y_k \leq U_k | m) \geq \beta) \\
&= \Pr\left(\int_{\frac{1-F_\theta(U_k)}{F_\theta(U_k)}}^{\frac{2k}{2(m-k+1)}} f_{2(m-k+1),2k}(x)dx \geq \beta\right) \\
&= \Pr\left(\int_0^{\frac{1-F_\theta(U_k)}{F_\theta(U_k)}} \frac{2k}{2(m-k+1)} f_{2(m-k+1),2k}(x)dx \leq 1-\beta\right) \\
&= \Pr\left(\frac{1-F_\theta(U_k)}{F_\theta(U_k)} \frac{2k}{2(m-k+1)} \leq q_{2(m-k+1),2k;1-\beta}\right) \\
&= \Pr\left(F_\theta(U_k) \geq \frac{k}{k+(m-k+1)q_{2(m-k+1),2k;1-\beta}}\right) \\
&= \Pr\left(\exp\left(-\frac{U_k - \mu}{\sigma}\right) \leq \frac{(m-k+1)q_{2(m-k+1),2k;1-\beta}}{(m-k+1)q_{2(m-k+1),2k;1-\beta} + k}\right) \\
&= \Pr\left(\frac{U_k - \mu}{\sigma} \geq -\ln\left(\frac{(m-k+1)q_{2(m-k+1),2k;1-\beta}}{(m-k+1)q_{2(m-k+1),2k;1-\beta} + k}\right)\right) \\
&= \Pr(V \geq -\eta_U W - \ln \delta_{1-\beta}) = \int_{-\eta_U w - \ln \delta_{1-\beta}}^\infty h(v)dv, \quad (30)
\end{aligned}$$

where

$$\eta_U = \frac{L_k - X_1}{S_1}, \quad (31)$$

is the upper tolerance factor,

$$\delta_{1-\beta} = \frac{(m-k+1)q_{2(m-k+1),2k;1-\beta}}{(m-k+1)q_{2(m-k+1),2k;1-\beta} + k}. \tag{32}$$

It follows from (18), (27) and (30) that

$$\begin{aligned} \eta_U &= \arg \left( \int_0^\infty \int_{-\eta_U w - \ln \delta_{1-\beta}}^\infty h(v)g(w)dv dw = \gamma \right) \\ &= \frac{1}{n} \left[ 1 - \left( \frac{\delta_{1-\beta}^n}{\gamma} \right)^{1/(r-1)} \right]. \end{aligned} \tag{33}$$

Taking into account (31) and (33), we have (28). This completes the proof.

*Corollary 3.1.* If  $k=m=1$ , then

$$U_k = \begin{cases} X_1 + \frac{S_1}{n} \left[ 1 - \left( \frac{(1-\beta)^n}{\gamma} \right)^{\frac{1}{r-1}} \right], & \text{if } n \geq \frac{\ln \gamma}{\ln(1-\beta)}, \\ X_1 - \frac{S_1}{n} \left[ \left( \frac{(1-\beta)^n}{\gamma} \right)^{\frac{1}{r-1}} - 1 \right], & \text{if } n < \frac{\ln \gamma}{\ln(1-\beta)}. \end{cases} \tag{34}$$

The result similar to that of (34) can be found in [7].

*Remark 1.* It will be noted that an upper tolerance limit may be obtained from a lower tolerance limit by replacing  $\beta$  by  $1-\beta$ ,  $\gamma$  by  $1-\gamma$ .

### 4. Practical Example

An industrial firm has the policy to replace a certain device, used at several locations in its plant, at the end of 24-month intervals. It doesn't want too many of these items to fail before being replaced. Shipments of a lot of devices are made to each of three firms. Each firm selects a random sample of  $l=5$  items and accepts his shipment if no failures occur before a specified lifetime has accumulated. In order to find this specified lifetime, the manufacturer wishes to take a random sample of size  $n=15$  and to calculate the lower one-sided simultaneous tolerance limit  $L_{k=1}(S)$  which is expected to capture a certain proportion  $\beta=0.95$  or more of the population of selected items ( $m=3l$ ), with a given confidence level  $\gamma=0.95$ . This tolerance limit is such that one can say with a certain confidence  $\gamma$  that at least  $100\beta\%$  of the product selected for testing by firms will operate longer than  $L_1(S)$ . The resulting lifetimes (rounded off to the nearest month) of the initial sample of size  $n=15$  from a population of the aforementioned devices are given in Table 1.

**Table 1.** The resulting lifetimes of the initial sample of size  $n=15$ .

Observations (in terms of month intervals)							
$X_1$	$X_2$	$X_3$	$X_4$	$X_5$	$X_6$	$X_7$	$X_8$
8	9	10	12	14	17	20	25
$X_9$	$X_{10}$	$X_{11}$	$X_{12}$	$X_{13}$	$X_{14}$	$X_{15}$	
29	30	35	40	47	54	62	

Goodness-of-fit testing. It is assumed that

$$X_i \sim f_\theta(x) = \frac{1}{\sigma} \exp\left(-\frac{x-\mu}{\sigma}\right), \quad x \geq \mu, \tag{35}$$

where the parameters  $\mu$  and  $\sigma$  are unknown. Thus, for this example, we have that  $r=n=15$ ,  $m=3l=15$ ,  $k=1$ ,  $\beta=0.95$ ,  $\gamma=0.95$ ,

$$S = \left( X_1 = 8, \quad S_1 = \sum_{i=1}^n (X_i - X_1) = 266 \right). \tag{36}$$

It can be shown that the

$${}^jU(0,1) = 1 - \frac{\left( \sum_{i=2}^{j+1} (n-i+1)(X_i - X_{i-1}) \right)^j}{\sum_{i=2}^{j+2} (n-i+1)(X_i - X_{i-1})}, \quad j = 1(1)n-2, \tag{37}$$

are i.i.d.  $U(0,1)$  rv's (Nechval *et al.* [11]). We assess the statistical significance of departures from the two-parameter exponential model (10) by performing the Kolmogorov-Smirnov goodness-of-fit test. We use the  $K$  statistic (Muller *et al.* [12]). The rejection region for the  $\alpha=0.05$  level of significance is  $\{K \geq K_{n,\alpha}\}$ . The percentage points for  $K_{n,\alpha}$  were given by Muller *et al.* [12]. For this example,

$$K = 0.280 < K_{n=13,\alpha=0.05} = 0.361. \tag{38}$$

Thus, there is not evidence to rule out the two-parameter exponential model.

Now the lower one-sided simultaneous  $\beta$ -content tolerance limit at the confidence level  $\gamma$ ,  $L_1 \equiv L_1(S)$  (on the order statistic  $Y_1$  from a set of  $m$  future ordered observations  $Y_1 \leq \dots \leq Y_m$ ) can be obtained from (20). Since

$$n = 15 < \frac{\ln(1-\gamma)}{\ln \delta_\beta} = 292, \tag{39}$$

where

$$\delta_\beta = \frac{(m-k+1)q_{2(m-k+1),2k;\beta}}{(m-k+1)q_{2(m-k+1),2k;\beta} + k} = 0.989848, \tag{40}$$

it follows from (20) that

$$L_1(S) = X_1 - \frac{S_1}{n} \left[ \left( \frac{\delta_\beta^n}{1-\gamma} \right)^{\frac{1}{r-1}} - 1 \right] = 4. \tag{41}$$

*Statistical inference.* Thus, the manufacturer has 95% assurance that no failures will occur in the proportion  $\beta=0.95$  or more of the population of selected items before  $L_1 = 4$  month intervals.

## 5. Conclusion

Tolerance limits enjoy a fairly rich history in the literature and have a very important role in engineering and manufacturing applications. In contrast to other statistical limits commonly used for statistical inference, the tolerance limits (especially for the order statistics) are used relatively rarely. One reason is that the theoretical concept and computational complexity of the tolerance limits is significantly more difficult than that of the standard confidence and prediction limits. Thus it becomes necessary to use the innovative approaches which will allow one to construct tolerance limits on future order statistics for many populations.

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